

A Multi-marker Test for Analyzing Paired Transplant Genetic Data Victoria L Arthur^{1,} Zhengbang Li^{1,2}, Rui Cao³, Marylyn D. Ritchie⁴, Weihua Guan³, and Jinbo Chen¹

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Introduction

Evidence suggests that donor/recipient (D/R) matching in some genetic regions may impact transplant outcomes^{1,2}. Most available matching scores account for single-nucleotide polymorphism (SNP) matching only or matching across a long range of different gene regions, making it hard to interpret association findings. In this work, we propose a multi-marker method, the Joint Score Test (JST), to jointly test for association between R genotype SNP effects and a gene-based matching score with transplant outcome. Additionally, we use a penalized testing method to test for association of a gene-based matching score with transplant outcome while adjusting for possible R genotype SNP effects.

Model and Notation

GLM for outcome Y_i (i = 1, ..., n):

$$g(\mu) = \alpha_0 + W_i \alpha + X_i \beta + Z_i \gamma$$

- g(): link function
- $\mu = E(Y)$
- $W_i = (W_{i1}, ..., W_{iK})$: vector of K covariates for D/R pair *i*
- $X_i = (X_{i1}^R, ..., X_{im}^R)$: R genotype vector of *m* SNPs for recipient i
- Z_i : single, gene-based genetic matching score value for D/R pair *i*

Null hypotheses of interest:

$$H_0: \boldsymbol{\beta} = 0 \text{ and } \gamma = 0$$

and
$$\gamma = 0$$

Gene Based Scores:

where $D(X_{ij}^{D}, X_{ij}^{R})$ is a measured distance between the D and R genomes **Distance Measures:**

 D_M

Joint Score Test (JST)

logit P

Methods

 $Z_i = \sum_{j=1}^m D(X_{ij}^D, X_{ij}^R),$

Allogenomics Mismatch Score³

$$D_{AMS} = \sum_{a \in X_{ij}^{D}} \begin{cases} 0 \text{ if } a \in X_{ij}^{R} \\ 1 \text{ otherwise} \end{cases}$$

Where *a* denotes alleles of a genotype

Binary Mismatch Score⁴

$$M_{IM} = \begin{cases} 1 \text{ if } \exists \ a \ \in X_{ij}^D \text{ such that } a \ \notin X_{ij}^R \\ 0 \text{ otherwise} \end{cases}$$

IBS Mismatch Score

$$D_{IBS} = |X_{ij}^D - X_{ij}^R|$$

Incompatibility Score

 $D_{Incomp} = \begin{cases} 1 \text{ if } X_{ij}^D \neq X_{ij}^R \\ 0 \text{ otherwise} \end{cases}$

 $\hat{p}_1(\boldsymbol{W}_i) \equiv p(Y_i = 1 | \boldsymbol{W}_i; \hat{\alpha}_0; \hat{\boldsymbol{\alpha}})$: predicted probability of $Y_i = 1$ based on the null model:

logit
$$Pr(Y_i = 1) = \alpha_0 + \sum_{k=1}^{K} W_{ik} \alpha_k \equiv \alpha_0 + W_i \alpha$$

 $\hat{\alpha}_0$ and $\hat{\alpha}$: maximum likelihood estimates of α_0 and

• \hat{X}_{ij} : fitted value from $X_{ij} = \theta_0 + \sum_{k=1}^{K} W_{ik} \theta_k$ • \hat{Z}_i : fitted value from $Z_i = \tau_0 + \sum_{k=1}^{K} W_{ik} \tau_k$ Weights for above models are $\hat{p}_1(W_i)$ {1 – $\hat{p}_1(W_i)$ for R *i* or D/R pair *i*

- Define $B_i = (X_i, Z_i)$ and $\widehat{B}_i = (\widehat{X}_i, \widehat{Z}_i)$
- matching score
- Construct Hotelling's T^2 statistic as

•
$$V = \begin{bmatrix} V^R & C^{RS} \\ C^{SR} & V^S \end{bmatrix} = \begin{bmatrix} C \\ C \end{bmatrix}$$

- JST is based on Eigen decomposition of V^R

- JST is constructed as

$$\begin{pmatrix} \boldsymbol{U}^{\boldsymbol{P}\boldsymbol{R}} \\ \boldsymbol{U}^{\boldsymbol{S}} \end{pmatrix}^{T} \begin{bmatrix} \boldsymbol{I}_{\boldsymbol{S}\times\boldsymbol{S}} \\ \boldsymbol{C}\boldsymbol{o}\boldsymbol{v}(\boldsymbol{U}^{\boldsymbol{S}}, \boldsymbol{U}^{\boldsymbol{P}\boldsymbol{R}}) \end{bmatrix}$$

• JST is asymptotically distributed as χ^2_{s+1}

Penalized Score Test⁵

- Define $X^* = \{1, W, X, Z\}, n \times p$ matrix, (p = k + m + 2)
- $\boldsymbol{\omega} = \{\alpha_0, \boldsymbol{\alpha}, \boldsymbol{\beta}, \boldsymbol{\gamma}\}, p$ -dimensional vector
- PDF of Y in exponential form: $\exp\left(\frac{Y_i X_i^* \omega}{\omega}\right)$
- General null hypothesis: Cω
- Our null hypothesis: $\omega_{0,M}$



Methods II

• $U = (B - \widehat{B}) \{Y - \widehat{p}_1\}$, where U is the vector of likelihood score statistics for all R SNPs and the • U is asymptotically distributed as $\mathcal{N}_{m+1}(0, V)$ $nU'\widehat{V}^{-1}U \sim \chi^2_{m+1}$ • Can improve power for large *m* by eliminating \hat{V}^{-1} $Var(\boldsymbol{U}^{\boldsymbol{R}}) \quad Cov(\boldsymbol{U}^{\boldsymbol{R}}, U^{\boldsymbol{S}})$ $Cov(U^S, U^R)$ $Var(U^{S})$ • $A = [a_1, a_2, ..., a_m]$: $m \times m$ matrix of eigenvectors of \widehat{V}^{R} with eigenvalues $(\lambda_{1}, \lambda_{2}, ..., \lambda_{m}), \lambda_{1} \geq \cdots \geq \lambda_{m}$ Extract first s < m PCs, $A_s = [a_1, a_2, ..., a_s]$ • Define U^{PR} : vector of $U^{R'}a_l/\sqrt{\lambda_l}$, l = 1, 2, ..., s $\begin{bmatrix} Cov(\boldsymbol{U}^{\boldsymbol{P}\boldsymbol{R}}, U^{\boldsymbol{S}}) \\ Var(U^{\boldsymbol{S}}) \end{bmatrix}^{-1} \begin{pmatrix} \boldsymbol{U}^{\boldsymbol{P}\boldsymbol{R}} \\ U^{\boldsymbol{S}} \end{pmatrix}$

$$\frac{-b(\boldsymbol{X}_{\boldsymbol{i}}^*\boldsymbol{\omega})}{\phi_0} \bigg) c(Y)$$

$$\omega_{0,M} = t$$

 $\omega_{0,M} = 0$, where $\omega_{0,M} = 2$

Methods III

Partially penalized likelihood function:

$$L_n(\boldsymbol{\omega}, \boldsymbol{\lambda}) = \frac{1}{n} \sum_{i=1}^n \{Y_i \boldsymbol{X}_i^* \boldsymbol{\omega} - b(\boldsymbol{X}_i^* \boldsymbol{\omega})\} -$$

- p_{λ} (): penalty function with tuning parameter λ
- Estimates of $\boldsymbol{\omega}$ under H_0 and H_a :

$$\widehat{\boldsymbol{\omega}}_{\mathbf{0}} = \arg \max_{\boldsymbol{\omega}} L_n(\boldsymbol{\omega}, \lambda_{n,0}) \quad \text{subject}$$
$$\widehat{\boldsymbol{\omega}}_{a} = \arg \max_{\boldsymbol{\omega}} L_n(\boldsymbol{\omega}, \lambda_{n,a}).$$

- Forced penalties for $\{\alpha_0, \alpha\}$ to be 0 so only elements of β were penalized
- Penalized score test statistic (T_s)

$$\{\boldsymbol{Y} - \boldsymbol{\mu}(\boldsymbol{X}^* \widehat{\boldsymbol{\omega}}_{\mathbf{0}})\}^T \begin{pmatrix} \boldsymbol{X}^* _M \\ \boldsymbol{X}^* _{\widehat{S}_0} \end{pmatrix} \widehat{\boldsymbol{\Omega}}_0 \begin{pmatrix} \boldsymbol{X}^* _M \\ \boldsymbol{X}^* _{\widehat{S}_0} \end{pmatrix}^T \{\boldsymbol{Y} - \widehat{\boldsymbol{S}}_0 = \{j \in M^C : \widehat{\boldsymbol{\omega}}_{0,j} \neq 0\}$$

• $\widehat{\boldsymbol{\Omega}}_0 = \begin{pmatrix} \boldsymbol{X}_M^{*T} \boldsymbol{\Sigma} (\boldsymbol{X}^* \widehat{\boldsymbol{\omega}}_{\mathbf{0}}) \boldsymbol{X}_M^* & \boldsymbol{X}_M^{*T} \boldsymbol{\Sigma} (\boldsymbol{X}^* \widehat{\boldsymbol{\omega}}_{\mathbf{0}}) \end{pmatrix}$

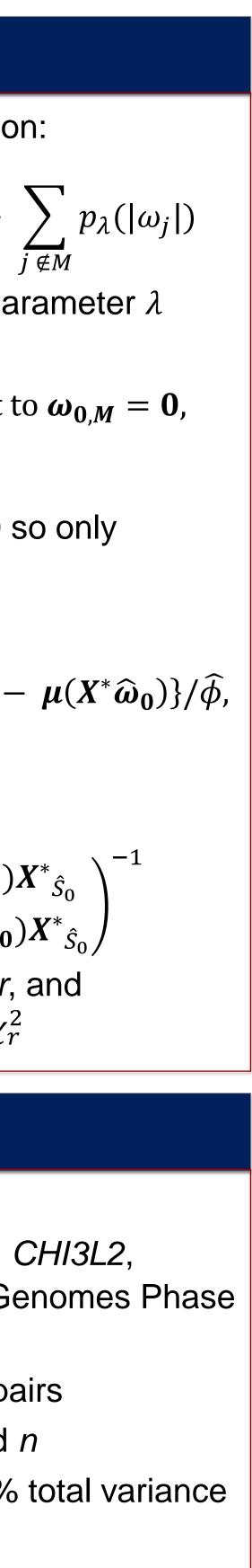
$$n \begin{pmatrix} \mathbf{X}_M \mathbf{Z} (\mathbf{X} \mathbf{\omega}_0) \mathbf{X}_M & \mathbf{X}_M \mathbf{Z} (\mathbf{X} \mathbf{\omega}_0) \\ \mathbf{X}^{*T}_{\hat{S}_0} \mathbf{\Sigma} (\mathbf{X}^* \widehat{\boldsymbol{\omega}}_0) \mathbf{X}_M^* & \mathbf{X}^{*T}_{\hat{S}_0} \mathbf{\Sigma} (\mathbf{X}^* \widehat{\boldsymbol{\omega}}_0) \end{pmatrix}$$

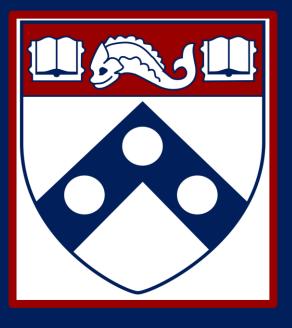
For a fixed number of constraints, r, and consistent estimator $\hat{\phi}$ for ϕ_0 , $T_S \sim \chi_r^2$

Simulations

Study Design

- Datasets for 3 gene regions (*NAT2*, *CHI3L2*, ASAH1) were sampled from 1000 Genomes Phase 3 reference using HapGen2⁶
- Sample size: n = 500 or 1000 D/R pairs
- 5000 simulations for each gene and *n*
- s values account for 85, 90, 95, 99% total variance explained by PCs





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Simulations Continued

- Options for power analyses:
 - 5, 15, 25% R genotype SNPs associated with outcome, Y
 - 5, 15, 25, 50, 75, 100% D/R matching associated with outcome, Y
 - Associated SNPs in low or high LD
 - Small (1.25), Medium (1.50), Large (2.00) OR per SNP or matching score
 - Outcome prevalence of 5, 10, 20%
- Compared JST to:
 - Standard GLM
 - SKAT
 - Penalized score test

Simulation Results

Method	Score	Prev. 20	Prev. 10	Prev. 5	Cont.	
	IBS	0.05	0.06	0.04	0.05	
JST	Incompatibility	0.05	0.05	0.05	0.05	
(s = 85%)	AMS	0.05	0.06	0.05	0.05	
,	Binary MM	0.05	0.06	0.05	0.05	
SKAT*	IBS	0.05	0.05	0.05	0.05	
	Incompatibility	0.05	0.05	0.05	0.05	
	AMS	0.05	0.05	0.05	0.05	
	Binary MM	0.05	0.05	0.05	0.05	
GLM	IBS	0.22	0.17	0.16	0.05	
	Incompatibility	0.22	0.15	0.15	0.05	
GLIVI	AMS	0.22	0.17	0.16	0.05	
	Binary MM	0.23	0.19	0.15	0.06	
	IBS	0.05	0.07	0.11	0.08	
Pen.	Incompatibility	0.05	0.08	0.11	0.08	
Score	AMS	0.05	0.09	0.11	0.09	
	Binary MM	0.05	0.08	0.10	0.10	

Table 1: Results of Type I Error simulations for JST using the gene NAT2
 with 500 D/R pairs. Score refers to which score was fit as Z_i . Results were similar for JST with *s* values of 90, 95, 99% variance explained. *SKAT was fit using an unweighted linear kernel.

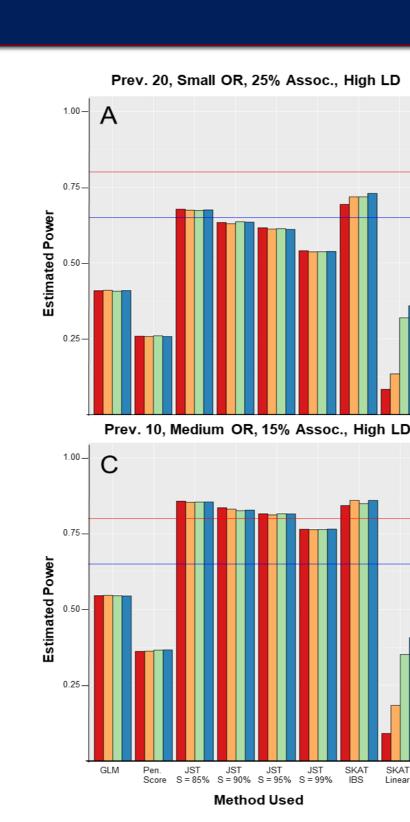
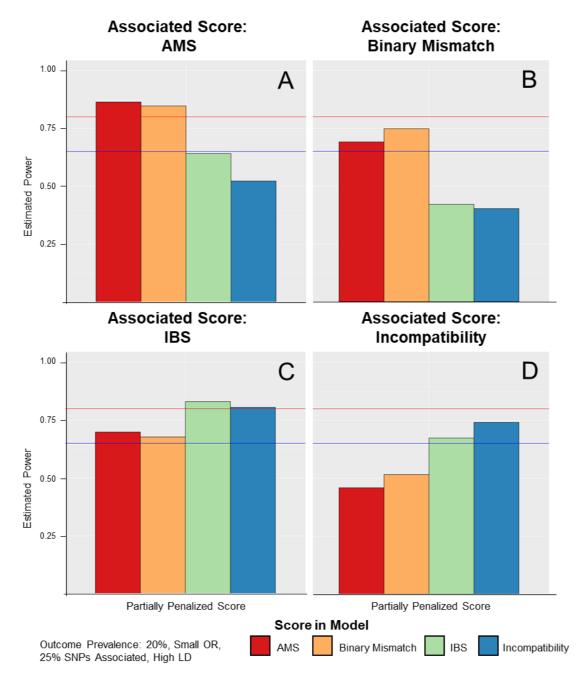
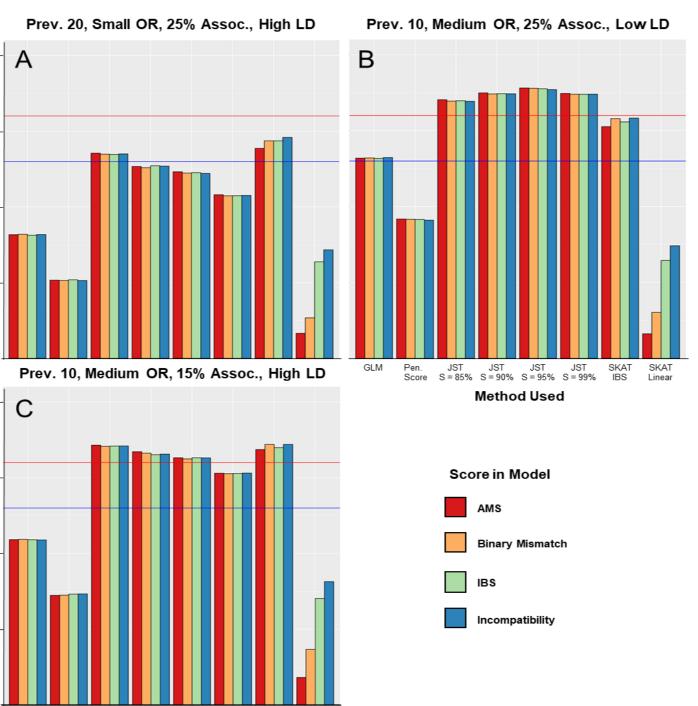


Figure 1: Power estimates from simulations using the gene NAT2 and 1000 pairs of donors and recipients under the scenario that recipient genotype SNPs were associated with outcome. The horizontal blue line corresponds to 65% power and the horizontal red line corresponds to 80% power.



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Simulation Results Continued



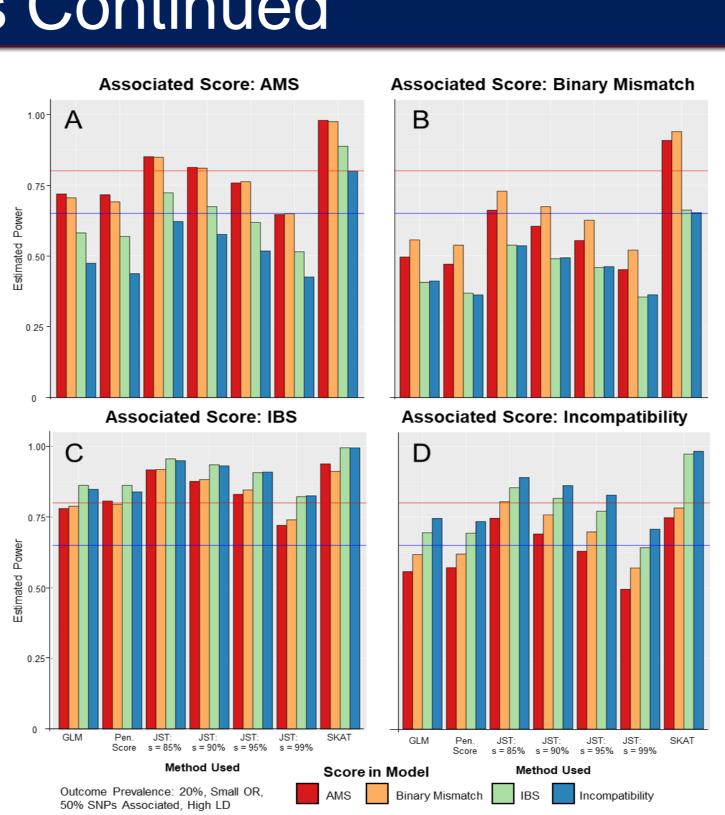
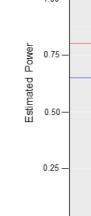


Figure 2: Power estimates from simulations using the gene NAT2 and 1000 pairs of donors and recipients under the scenario that the genebased score was associated with outcome. The horizontal blue line corresponds to 65% power and the horizontal red line corresponds to 80% power.

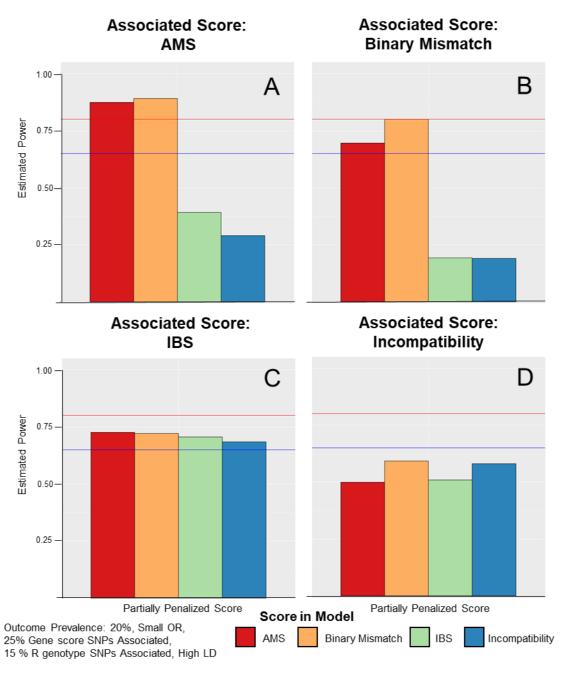
Figure 3 (Left): Estimated power plots for simulations testing whether gene-based score was associated with outcome. All models were fit using the partially penalized score test, with 1000 donor/recipient pairs. The blue line corresponds to 65% power and the horizontal red line corresponds to 80% power.

Figure 4 (Right): Estimated power plots for simulations testing whether gene-based score was associated with outcome. All models were fit using the partially penalized score test, with 1000 donor/recipient pairs. The blue line corresponds to 65% power and the horizontal red line corresponds to 80% power.









Real Data Analy

- Samples: 404 D/R kidney transplant (56 cases of Acute Rejection)
- Genome-wide SNPs (785,458 Bi-alle
- Grouped by 25,265 genes (physical

		JST Results	
Gene ID	IBS Score	P-value	AMS Sc
AC119677.1	29.25	4.46E-07	13.08
OVCH2	33.14	1.12E-06	28.72
	SKAT Results		
Gene ID	IBS Score	P-value	AMS Sc
OVCH2	454.93	7.916E-06	230.0
AC119677.1	107.41	5.143E-04	32.94
	Matchin	g Score Test	Results
Gene ID	IBS Score	P-value	AMS Sc
OVCH2	19.31	1.11E-05	12.01
AC119677.1	16.19	5.74E-05	4.96

Table 2: After Bonferroni correction, two genes were for in joint testing. Of these, OVCH2 was also found to be SKAT testing and the matching score only test. Results score and binary mismatch score match those for the r

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