

A Multi-marker Test for Analyzing Paired Transplant Genetic Data Victoria L Arthur1, Zhengbang Li 1,2 , Rui Cao³ , Marylyn D. Ritchie 4 , Weihua Guan³ , and Jinbo Chen¹

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Introduction

Model and Notation

GLM for outcome Y_i $(i = 1, ..., n)$:

Methods Methods II Methods III

Simulations

- $g()$: link function
- $\mu = E(Y)$
- $W_i = (W_{i1}, ..., W_{iK})$: vector of *K* covariates for D/R pair *i*
- $X_i = (X_{i1}^R, ..., X_{im}^R)$: R genotype vector of *m* SNPs for recipient *i*
- \bullet Z_i : single, gene-based genetic matching score value for D/R pair *i*

where $D(X_{ij}^D,X_{ij}^R)$ is a measured distance between the D and R genomes **Distance Measures:**

 D_M

Evidence suggests that donor/recipient (D/R) matching in some genetic regions may impact transplant outcomes^{1,2}. Most available matching scores account for single-nucleotide polymorphism (SNP) matching only or matching across a long range of different gene regions, making it hard to interpret association findings. In this work, we propose a multi-marker method, the Joint Score Test (JST), to jointly test for association between R genotype SNP effects and a gene-based matching score with transplant outcome. Additionally, we use a penalized testing method to test for association of a gene-based matching score with transplant outcome while adjusting for possible R genotype SNP effects.

> $D_{Incomp} = \{$ 1 if $X_{ij}^D \neq X_{ij}^R$ 0 otherwise

• $\hat{p}_1(\boldsymbol{W}_i) \equiv p(Y_i = 1 | \boldsymbol{W}_i; \hat{\alpha}_0; \hat{\boldsymbol{\alpha}})$: predicted probability of $Y_i = 1$ based on the null model:

$$
g(\mu) = \alpha_0 + W_i \alpha + X_i \beta + Z_i \gamma
$$

- Define $\boldsymbol{B}_i = (X_i, Z_i)$ and $\widehat{\boldsymbol{B}}_i = (\widehat{X}_i)$
- matching score
-
- Construct Hotelling's *T*² statistic as
-

Null hypotheses of interest:

$$
H_0: \beta = 0 \text{ and } \gamma = 0
$$

and

$$
\gamma = 0
$$

Gene Based Scores:

- JST is based on Eigen decomposition of V^R
-
- Extract first $s < m$ PCs, $A_s = [a_1, a_2, ..., a_s]$
-
- JST is constructed as

Allogenomics Mismatch Score 3

$$
D_{AMS} = \sum_{a \in X_{ij}^D} \begin{cases} 0 \text{ if } a \in X_{ij}^R \\ 1 \text{ otherwise} \end{cases}
$$

Where *a* denotes alleles of a genotype

Binary Mismatch Score 4

$$
M_{IM} = \begin{cases} 1 \text{ if } \exists a \in X_{ij}^D \text{ such that } a \notin X_{ij}^R\\ 0 \text{ otherwise} \end{cases}
$$

IBS Mismatch Score

$$
D_{IBS} = |X_{ij}^D - X_{ij}^R|
$$

Incompatibility Score

Joint Score Test (JST)

 $logit$ F

- α
-
- \hat{Z}
-

 $Z_i = \sum_{j=1}^m D(X_{ij}^D, X_{ij}^R),$

• For a fixed number of constraints, *r*, and consistent estimator $\widehat{\phi}$ for ϕ_0 , $T_S{\sim}\chi^2_r$

$$
Pr(Y_i = 1) = \alpha_0 + \sum_{k=1}^{K} W_{ik} \alpha_k \equiv \alpha_0 + W_i \alpha
$$

$$
\hat{\alpha}
$$
 maximum likelihood estimates of α_0 are

• $\hat{\alpha}_0$ and $\hat{\alpha}$: maximum likelihood estimates of α_0 and

• \hat{X}_{ij} : fitted value from $X_{ij} = \theta_0 + \sum_{k=1}^K W_{ik} \theta_k$ *i*: fitted value from $Z_i = \tau_0 + \sum_{k=1}^K W_{ik} \tau_k$ • Weights for above models are $\hat{p}_1(\boldsymbol{W}_i)\{1-\}$ $\hat{p}_1(\pmb{W}_{\pmb{i}}) \}$ for R *i* or D/R pair *i*

•
$$
V = \begin{bmatrix} V^R & C^{RS} \\ C^{SR} & V^S \end{bmatrix} = \begin{bmatrix} C & C & C \\ C & C & C \end{bmatrix}
$$

$$
\begin{pmatrix}\n\boldsymbol{U}^{PR} \\
\boldsymbol{U}^S\n\end{pmatrix}^T\n\begin{bmatrix}\nI_{S\times S} & Cov(\boldsymbol{U}^{PR}, \boldsymbol{U}^S) \\
Cov(\boldsymbol{U}^S, \boldsymbol{U}^{PR}) & Var(\boldsymbol{U}^S)\n\end{bmatrix}^{-1}\n\begin{pmatrix}\n\boldsymbol{U}^{PR} \\
\boldsymbol{U}^S\n\end{pmatrix}
$$
\n• JST is asymptotically distributed as χ^2_{s+1}

Penalized Score Test 5

- Define $X^* = \{1, W, X, Z\}, n \times p$ matrix, $(p = k + m + 2)$
- $\boldsymbol{\omega} = {\alpha_0, \boldsymbol{\alpha}, \boldsymbol{\beta}, \gamma}$, *p*-dimensional vector
- PDF of *Y* in exponential form: exp $Y_i X_i^* \omega - b(X_i^*$
- General null hypothes $C\omega$
- Our null hypothesis: $\omega_{0,M}$

 $_i$, \hat{Z}_i) • $U = (B - \widehat{B})\{Y - \widehat{p}_1\}$, where U is the vector of likelihood score statistics for all R SNPs and the • *U* is asymptotically distributed as \mathcal{N}_{m+1} (0, V) $n\bm{U}' \widehat{\bm{V}}^{-1}\bm{U} \sim \bm{\chi}^{\bm{2}}_{\bm{m+1}}$ • Can improve power for large m by eliminating \widehat{V}^{-1} $Var(U^R)$ $Cov(U^R, U^S)$ $Cov(U^S, U)$ R) $Var(U^S)$ • $A = [a_1, a_2, ..., a_m]$: $m \times m$ matrix of eigenvectors of \widehat{V}^R with eigenvalues $(\lambda_1, \lambda_2, ..., \lambda_m)$, $\lambda_1 \geq \cdots \geq \lambda_m$ • Define \boldsymbol{U}^{PR} : vector of $\boldsymbol{U}^{R'}a_l/\sqrt{\lambda_l}$, $l=1,2,...,s$

$$
\left(\frac{-b(X_i^*\omega)}{\phi_0}\right)c(Y)
$$

$$
u_{0,M} = t
$$
\n
$$
u_{0,M} = 0 \text{ where } u_{0,M} = 1
$$

• Partially penalized likelihood function:

$$
L_n(\boldsymbol{\omega}, \lambda) = \frac{1}{n} \sum_{i=1}^n \{ Y_i \boldsymbol{X}_i^* \boldsymbol{\omega} - b(\boldsymbol{X}_i^* \boldsymbol{\omega}) \} -
$$

- $p_{\lambda}($: penalty function with tuning parameter λ
- Estimates of ω under H_0 and H_a :

$$
\widehat{\boldsymbol{\omega}}_{0} = \underset{\boldsymbol{\omega}}{\arg \max} L_{n}(\boldsymbol{\omega}, \lambda_{n,0}) \text{ subject}
$$

$$
\widehat{\boldsymbol{\omega}}_{a} = \underset{\boldsymbol{\omega}}{\arg \max} L_{n}(\boldsymbol{\omega}, \lambda_{n,a}).
$$

- Forced penalties for $\{\alpha_0, \alpha\}$ to be 0 so only elements of β were penalized
- Penalized score test statistic (T_S)

$$
\{Y - \mu(X^*\widehat{\omega}_0)\}^T \begin{pmatrix} X^*_{M} \\ X^*_{\widehat{S}_0} \end{pmatrix} \widehat{\Omega}_0 \begin{pmatrix} X^*_{M} \\ X^*_{\widehat{S}_0} \end{pmatrix}^T \{Y - \widehat{S}_0 = \{j \in M^C : \widehat{\omega}_{0,j} \neq 0\}
$$

• $\widehat{\Omega}_0 = \begin{pmatrix} X^*_{M} \Sigma(X^*\widehat{\omega}_0) X^*_{M} & X^*_{M} \Sigma(X^*\widehat{\omega}_0) \end{pmatrix}$

$$
n\begin{pmatrix} X_M^{*T} \Sigma (X^* \widehat{\omega}_0) X_M^* & X_M^{*T} \Sigma (X^* \widehat{\omega}_0) \\ X^{*T} \widehat{S}_0 \Sigma (X^* \widehat{\omega}_0) X_M^* & X^{*T} \widehat{S}_0 \Sigma (X^* \widehat{\omega}_0) \end{pmatrix}
$$

Study Design

- Datasets for 3 gene regions (*NAT2*, *CHI3L2*, ASAH1) were sampled from 1000 Genomes Phase 3 reference using HapGen2 6
- Sample size: $n = 500$ or 1000 D/R pairs
- 5000 simulations for each gene and *n*
- *s* values account for 85, 90, 95, 99% total variance explained by PCs

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References

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- Options for power analyses:
	- 5, 15, 25% R genotype SNPs associated with outcome, *Y*
	- 5, 15, 25, 50, 75, 100% D/R matching associated with outcome, *Y*
	- Associated SNPs in low or high LD
	- Small (1.25), Medium (1.50), Large (2.00) OR per SNP or matching score
	- Outcome prevalence of 5, 10, 20%
- Compared JST to:
	- Standard GLM
	- SKAT
	- Penalized score test

Simulation Results

Table 1: Results of Type I Error simulations for JST using the gene *NAT2* with 500 D/R pairs. Score refers to which score was fit as *Zⁱ* . Results were similar for JST with *s* values of 90, 95, 99% variance explained. *SKAT was fit using an unweighted linear kernel.

Figure 1: Power estimates from simulations using the gene NAT2 and 1000 pairs of donors and recipients under the scenario that recipient genotype SNPs were associated with outcome. The horizontal blue line corresponds to 65% power and the horizontal red line corresponds to 80% power.

Method User

Simulations Continued Number 2014 Simulation Results Continued Real Data Analy

Figure 2: Power estimates from simulations using the gene *NAT2* and 1000 pairs of donors and recipients under the scenario that the genebased score was associated with outcome. The horizontal blue line corresponds to 65% power and the horizontal red line corresponds to 80% power.

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Figure 3 (Left): Estimated power plots for simulations testing whether gene-based score was associated with outcome. All models were fit using the partially penalized score test, with 1000 donor/recipient pairs. The blue line corresponds to 65% power and the horizontal red line corresponds to 80% power.

Figure 4 (Right): Estimated power plots for simulations testing whether gene-based score was associated with outcome. All models were fit using the partially penalized score test, with 1000 donor/recipient pairs. The blue line corresponds to 65% power and the horizontal red line corresponds to 80% power.

- Samples: 404 D/R kidney transplant (56 cases of Acute Rejection)
- Genome-wide SNPs (785,458 Bi-alle
- Grouped by 25,265 genes (physical

Table 2: After Bonferroni correction, two genes were fore in joint testing. Of these, $OVCH2$ was also found to be SKAT testing and the matching score only test. Results score and binary mismatch score match those for the nonbinary mismatch score